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From one to many (and back again): an alternating renewal process describes the buildup of perceptual grouping Sara Steele 1,∗, Daniel Tranchina 2,3 and John Rinzel 1,2

Research Article 17 September 2013

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Research Topic

ABSTRACT

Previous reports of buildup describe how perceptual organization becomes more complex over time, in that different stimulus features are gradually more likely to be grouped into different perceptual representations. A common interpretation of this finding is that there must be some underlying mechanism of accumulation, for instance of sensory evidence (informational) or adaptation (physiological). We present a theoretical framework– an alternating renewal process– that shows that even in the absence of any accumulative process, such buildup can be observed in a system in which the observer takes random samples from alternative perceptual organizations, as in perceptual bistability. Using this theory, we can describe the short term dynamics of buildup observed on short trials in terms of the long term statistics of percept durations for each alternative perceptual organization. This model performs well in describing buildup functions computed from pseudo-neuronal simulations of neural populations with competition architecture that undergo alternations. Even when these competition model simulations produce history dependence through accumulation of adaptation, our alternation model, which neglects history dependence from accumulative processes, is able to predict the buildup function. This indicates that accumulation is not a necessary feature to produce or to describe a buildup function. Rather, any model producing buildup will have an underlying alternating renewal process, the statistics of which are sufficient to account for buildup.

Keywords: stream segregation; alternating renewal process; bistable perception; perceptual dynamics; perceptual organization

1 INTRODUCTION

For some stimuli in the auditory and visual modalities with ambiguous grouping cues, the probability of perceptual “splitting” increases with time. One example of this buildup is found in auditory stream segregation. A number of studies using ambiguous ABA- tone sequences (?) have shown that perceptual splitting of sound events with different acoustic features requires time to build up (??). A and B refer to tones at different frequencies, and - represents a silent interval. Depending on the difference in frequency and presentation rate of the tones, listeners might be more likely to hear grouped triplet patterns integrated into a galloping rhythm, or segregated streams of tones at separate frequencies (Figure ??). There is typically a period of buildup over which the probability of the segregated percept increases, starting from the initiation of a presentation (??) or a switch in the focus of attention (?). A similar phenomenon has also been reported in the visual modality. When viewing ambiguous moving plaids constructed from moving square wave gratings at intermediate speed and angle, observers have reported first experiencing coherent motion of a unified plaid pattern, even when, in the long term, their perception is biased towards transparent motion of the individual gratings (?). The psychophysical observation of this buildup in probability for split perceptual organization is accompanied by reports of perceptual alternations (?)That is, because of the ambiguity of the stimuli, subjects report experiencing sudden switches between one perceptual state and the other (??) over long presentations, even after the buildup period.

Buildup in the auditory domain has been a subject of great interest because it is thought to be a product of the same mechanisms that enable listeners to perceptually group components of an acoustic signal according to the sources that produced them, a process known as stream segregation. Many explanations of buildup appeal to proposed mechanisms of stream segregation. One theoretical explanation for the perceptual organizations observed with the ABA- stimuli is grouping by coactivation (for a review, see ?). Sound signals that excite the same population of neurons are grouped, whereas those which activate separate populations are perceived as coming from separate sources, that is, split. Some theories (???) propose that the buildup of stream segregation reflects the accumulation of adaptation over seconds, or multi-second habituation, which leads to a decrease over time of activation by the other tone of auditory neurons tuned to each of the two tone frequencies.

The accumulation-based account of buildup can explain the perceptual switch from the grouped to the split percept; however, it is unable to explain why subjects undergo continued switches between percepts. We show that buildup can be described well by a model that uses the statistics of alternations, but no accumulation. The gradual increase in probability of a split percept over time could reflect the dynamics of an entirely random underlying alternating renewal process with a given initial state.

The long-term dynamics of perceptual bistability consist of alternations between mutually excusive percepts. The duration histogram of each percept is well-fit by a gamma density (??). We believe that the short-term buildup in probability of split percepts, observed when short trial perceptual timecourses are averaged, could reflect the dominance duration distributions observed over long trials. To examine this relationship we use the theoretical framework of an alternating renewal process (). In this investigation we show that these distributions of percept durations, without considering accumulation, can account for the experimentally observed dynamics of buildup for a stimulus with ambiguous grouping, as follows:

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fq

coherent

transparent

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1. the perceptual state alternates back and forth between grouped and split 2. the durations for these perceptual epochs are random, independent and stationary 3. the initial percept on a given trial is always the grouped percept

In addition, we use existing computational models of perceptual bistability to explore how different mechanisms of alternation affect the model’s behavior with regard to buildup.

2 RESULTS

ALTERNATING RENEWAL PROCESS

The process of constructing empirical buildup curves involves the averaging over many trials of the timecourse of a random binary state variable (see Figure ??a, blue lines). In our statistical model, the initial state (percept) is fixed, but the dwell time in this state is a random variable characterized by its probability density function. Subsequently the system switches randomly between two states, each of which has its own fixed duration distribution.

We initially tested this theory in Monte Carlo simulations by simply constructing in silico perceptual timecourses according to the above assumptions (see Figure ??, (a)). For a given simulated trial timecourse, we draw alternating random samples from each of two distributions– one corresponding to the grouped state durations, and the other to the split state durations. We draw the first sample from the distribution corresponding to the grouped state, the second from that corresponding to the split state, and continue drawing samples from each distribution in alternation until the sum of all the durations exceeds the length of a trial. These trial durations are converted into discretized timecourses by assigning a value of 0 to time intervals during which the state corresponds to a grouped percept, and assigning a value of 1 to the time intervals for which the percept state was split. In Monte Carlo simulations, we produce an arbitrarily large number of such trial timecourses, then take the average at each time point. This simple process, with no history dependence between successive samples, produces buildup-like curves.

We have also derived an analytical solution to this process using a population density approach similar to that used in ? to model stochastic gene expression. There are a number of advantages to characterizing the buildup function in this way. First, with an analytical solution relating the distributions of durations for grouped and split percepts to the buildup function, it is theoretically possible to interconvert between buildup functions and the statistics of the dominance durations for each percept. We have developed this solution into a statistical switching model to reconstruct the buildup function from four parameters- the parameters for the gamma densities for grouped and split percept durations. This theoretical solution coincides with the Monte Carlo simulation results (see Figure ??b). This is convenient, as the analytical solution is both computationally less expensive than iterative Monte Carlo simulations, and the solution is exact. Furthermore, in the alternating renewal process framework, there is no explicit mechanism of accumulation. Therefore, this formulation challenges the conventional understanding of buildup of sound source segregation as requiring a gradual accumulative process.

2.1 COMPETITION MODEL SIMULATIONS PRODUCE BUILDUP AS A CONSEQUENCE OF ALTERNATIONS

There has been a great deal of work done to understand the neural computations that could underlie perceptual switching with ambiguous stimuli . Previous investigations (?????) have used population firing rate models with competition architecture to model perceptual bistability. In these pseudoneuronal mutual inhibition models, there are separate populations whose firing rates represent the perceptual strength of each interpretation of the stimulus. They make inhibitory connections onto one another, so the population with the highest firing rate typically dominates the other. These models were originally developed to describe binocular rivalry, but have also been used to account for the psychophysical results of experiments

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Steele et al. Alternating renewal process describes buildup with ambiguous grouping– namely, moving plaids with coherent/transparent motion (???) and triplets

with streaming integration/segregation (?).

In competition models, the relative firing rates of the two populations are taken to produce the simulated observer’s perceptual reports. The population with the higher firing rate corresponds to the dominant percept. Because the two populations mutually inhibit each other, in most cases only one population is active at any given time. In addition, each population undergoes adaptation in response to its own firing rate. The alternation of dominance epochs between the two populations can be driven by two mechanisms. If adaptation is strong enough, then the activity of the dominant population will decay over time, while the suppressed population recovers from any prior adaptation. This leads to periodic alternations between dominance states with oscillation dynamics. However, if adaptation is weak, the system will display attractor dynamics, in which alternations are driven by noise in the externally applied inputs. The brain appears to be a very noisy system, with random fluctuations occurring at multiple scales such as vesicular release and spiking variability.

(a)

0.2 0.15 0.1 0.05 0

0

0.08 0.06 0.04 0.02

T0(s) T1(s)

time (s)

Take alternating samples from duration distributions for grouped and split percepts to construct simulated perceptual timecourse for a trial

Assign a value of 0 to timepoints from grouped durations and a value of 1 to timepoints from split durations

Take the average at each timepoint over many trials to obtain the buildup function

Monte Carlo and Analytical Buildup Functions

(b)

time(s)

k = 1.45 θ = 3.26

1 0.9 0.8 0.7 0.6 0.5 0.4 0.3 0.2 0.1

10 20 30 40 50

grouped percept durations (s)

k = 2.08 θ = 5.12

00 0 10 20 30 40 50 0

split percept durations (s)

5 10 15

time (s)

The difference between oscillator

Figure 2: (a) Visualization of the alternating renewal process producing a buildup function. We used gamma probability density functions approximating duration distributions to construct Monte Carlo simulations. The analytical solution to the alternating renewal process model, shown as a red solid curve (bottom), allows us to predict buildup as a function of the distributions of dominance

t=0

trial 1

trial 2

. . .

trial n

durations of each of the perceptual states. (b) Monte Carlo simulation and attractor dynamics in these competition

models is most simply understood computed buildup function of 1000 trials (blue) approaches the by observing how the system would theoretical solution (red). Gamma distribution parameters chosen behave without noise (??). For oscillation from fits to a preliminary psychophysical dataset (not shown).

dynamics, adaptation would cause the dominant population to reduce its activity over time, reducing the inhibition on the suppressed population, allowing it to become active. In a noiseless system, stable fixed points in the system appear and disappear over time, and alternations will occur deterministically with a constant period. Noise in such a system will affect the distribution of dominance durations for each state, but is not required for switching. Conversely, attractor dynamics occurs when a system has multiple stable states at the same time. In the absence of noise, the initial conditions determine which state becomes active, and the system behaves in a winner-take-all fashion. That is, the population that becomes dominant first is permanently active, and the other population is permanently suppressed. However, injecting noise into such a system can cause switches from one stable state into another. In this case, the switching between perceptual dominance states would be caused by the noise itself.

We wanted to use the competition model as a test-bed for the theory of the alternating renewal process. We chose initial conditions that ensured that the grouped percept was always dominant at the beginning

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p(split)

probability "split"

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of the trial, as our hypothesis that alternations could produce buildup relies on this assumption. We did this in the simplest manner possible (see Figure ??), by setting the initial condition on the population representing the grouped percept to half its maximum value. All other initial conditions were set to 0. From the population firing rates we computed dominance durations, which were converted to binary timecourses and averaged over trials to produce competition model buildup functions. Using this method, we produced buildup functions for both oscillator and attractor competition models.

(a)

Input, I1 Adaptation, a1(t) a2

Competition model trial timecourse

u1 u2

0 5 10 15 20 γ = 0.1

To test the application of the alternating renewal process model, we viewed dominance durations of each state from the short simulated trials as samples from underlying gamma distributions. Because our competition model simulations generated trials that were only 20 seconds long, a large proportion of these durations were truncated by the end of the trial. We estimated gamma parameters k and θ that maximized the likelihood of the complete as well as the right-censored dominance durations for each model “perceptual” state. Using only the four parameters so obtained, we were able to get very strong fits to the buildup function (R-squared > 90%).

2.1.1 Psychophysically observed buildup is inconsistent with alternations under oscillation dynamics Previous work from our lab has shown that oscillation dynamics are inconsistent with a number of statistical features of the dominance durations reported in psychophysical experiments (?). The mean and

circular variance of dominance durations under this dynamical regime do not fall within the range of those observed for perceptual reports of ambiguous visual displays. Furthermore, when adaptation drives alternations in the dominance of population activity, we observe moderate and significant correlations between successive percepts. Data from the psychophysical literature suggests that the durations of subsequent percepts are only weakly correlated, if

at all (?).

u1

u2

(b)

1 0.5 0

500 grouped 0.7

400 1541 samples 300 k ≈ 2.14 0.6 200 θ ≈ 1.39 0.5 100

00 5 10 15 0.4 duration (s)

σ = 0.12 I1 = I2 =

0.6

600 400 200

split

0.3 1299 samples 0.2

k ≈ 2.3 θ ≈ 1.4 0.1

I2 r1 r2

00 0 5 10 15 0 2 4 6 8 10

duration (s) time (s)

Figure 3: (a) Mutual inhibition population firing rate model producing buildup. We choose initial conditions to ensure that the population representing the grouped percept, u1, is always dominant at the beginning of a given trial timecourse. (b) Competition model simulation results for parameters that produce attractor dynamics with noise-driven switching. Top, population activity timecourse for one 20-second trial. 500 trials were simulated to produce the buildup function, lower right (blue). Histograms of the dominance durations, with maximum likelihood estimated gamma density parameters, are shown in the lower left. These parameters allow us to compute analytically the resulting buildup function (red). The buildup function looks similar to those reported in the psychophysical literature, and the ARP prediction is good (R-Squared = 98%).

We wanted to examine buildup under each of these dynamical regimes in order to determine whether we could find correspondence between the buildup functions produced by competition models and those reported in the psychophysical literature. Previously reported psychophysical data suggest that buildup is largely a monotonic proess. To our knowledge, no

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p(split)

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(a)

1 underlying: 1.22 2.72, 3.53 3.29

5000 true: 1.22, 2.72 4000

0.9 0.8 0.7 0.6 0.5 0.4 0.3 0.2 0.1

(b) 1 0.9 0.8 0.7 0.6 0.5 0.4 0.3 0.2 0.1

LSE: 1.24 2.68, 3.33 3.72

00 10 20 30 40

2500 2000 1500 1000

true: 3.53, 3.29

fit: 3.33, 3.72

500 00 5 10 15 00 10 20 30 40 50

underlying: 2.68 2.65,1.06 3.64 4000 LSE: 2.12 4.46, 3.91 1.26 3000

2000 1000

duration (s)

true: 2.68, 2.65

fit: 2.12, 4.46

duration(s)

1000 00 5 10 15 00 10 20 30 40

duration(s)

Figure 5: We generated buildup functions, shown in blue, with Monte Carlo simulations from known gamma densities (shown in gray on the histograms). We then obtained estimates for those parameters by finding the least squares fit to the Monte Carlo simulated buildup functions from our four parameter analytical expression. The fitted buildup function and the gamma densities so recovered are shown in dashed red lines. (a) A successful recovery of the gamma density parameters from the buildup function. (b) One case in which the parameters that minimize squared error between the analytical and the Monte Carlo simulated buildup function do not match the underlying gamma densities used to produce it.

psychophysical experiments have shown a buildup timecourse with a damped oscillatory approach to steady state. Buildup functions produced under oscillation dynamics in our competition model (Figure ??), however, display damped oscillations reflecting the underlying periodicity of the mechanism of alternation. As previously reported ?, these buildup functions are derived from perceptual timecourses for which there are significant correlations between successive percept durations, such that the present perceptual state depends on the cumulative history of previous percepts. Although the correlation coefficient between durations of subsequent percepts was significant (r = 0.23), the fit to the buildup function by finding the parameters of density functions for long term percept durations was strong (R2 = 93%. Despite ignoring history dependence, the theory linking buildup to long-term dynamics gives accurate predictions.

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3000 2000 1000

fit: 1.24, 2.68

duration (s)

00 10 20 30 40 50

5000 4000 true: 1.06, 3.64

3000 fit: 3.91, 1.26 2000

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2.2

DURATION DISTRIBUTIONS ARE NOT WELL CONSTRAINED BY THE TRIAL AVERAGE

1 0.5

00

800 grouped 600 400 200

Oscillation dynamics timecourse

5 10 15 20

The resilience of the alternating renewal process to capture the dynamics of buildup in a variety of circumstances, including those with duration-to- duration correlations likely carries the burden of overflexibility– for any given buildup function, a wide range of pairs of distribution functions may provide a good fit. We generated buildup functions using Monte Carlo simulations as in Figure ?? and used our analytical solution to find the gamma density parameters that minimized the squared error between the Monte Carlo simulated and analytical buildup functions. In general, the parameters so found did not strongly fit the histograms of the duration samples used (see Figure ??). Specifically, we used one-sample Kolmogorov-Smirnov tests to determine whether the duration samples came from the gamma distribution obtained from the least squares fit to the buildup function. Only 24.5% of simulations produced good fits, and the average KS distance between sample and the fitted distribution was .09.

However, for the special case when the both perceptual states have identical duration distributions, the parameter estimation for the gamma density functions from the buildup function is much better. This circumstance occurs in competition model simulations when the inputs to the populations representing each percept are matched, e.g., when the stimulus is

2339 samples

k ≈ 5.98 θ ≈ 0.501

γ = 0.3 0.8 σ = 0.09

I1 = I2 = 0.6

00 2 4 6 8 10

0.7 0.6 0.5

800 600 400 200

duration(s) 0.4

splitgrouped 1984 samples

k ≈9 θ ≈ 0.365

0.3 0.2 0.1

00 2 4 6 8 10

duration (s)

00 2 4 6 8 10

time (s)

Figure 4: Competition model simulation of buildup under oscillation dynamics, for which switches between dominance states are driven by adaptation. Top, an example trial timecourse. The dominance durations are much more regularly timed, reflecting the clock- like periodicity of the underlying oscillator. These oscillations are dramatically present in the average. The prediction under the alternation hypothesis from the gamma parameters is still quite good (R-squared = 93%).

perfectly ambiguous. We generated buildup functions using Monte Carlo simulations from two identical gamma densities, and found the single pair of gamma parameters that minimized squared error between this and theoretical buildup function. For 1000 such simulations, 63% of the recovered parameters were indistinguishable from the sample empirical distribution by Kolmogorov-Smirnov tests, and in general the fits were much closer, with a mean KS distance of .04. This method of estimation may prove useful for providing at least a rough estimate of the duration distributions underlying steady state switching dynamics.

3 DISCUSSION

3.1 WHAT IF THE FIRST PERCEPT IS LONGER?

One issue we have not addressed so far in our presentation of the alternating renewal process model is inertia (?). For ambiguous displays, the time until the first perceptual switch is typically much longer than subsequent durations of the same percept. For stimuli with ambiguous grouping, the distribution of initial grouped percept durations is different from other grouped percepts. Our theoretical model is capable of computing the buildup function from both steady state and initial percept distributions; however, this would introduce a third duration distribution, and increase the number of parameters to 6. For simplicity’s

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p(split)

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sake, we have only shown the 4 parameter model, which assumes that the initial percept duration is drawn from the same distribution as other grouped percept durations.

Our competition model, as it stands, produces a different distribution for initial and subsequent grouped percepts. However, the initial percepts are shorter in mean duration. This can probably be fixed by using a different set of initial conditions to achieve the correct initial dominance state– for instance, by setting the initial conditions so that the other population, that which represents the split percept, is highly adapted at the beginning of the trial. However, such an approach would be arbitrary, and leaves something to be desired in terms of explanatory value. What would be better is a more detailed model that captures elements of the sensory coding and neural dynamics that produce the perceptual states in the first place. For instance, percept formation might be subserved by neural populations that respond only to particular spectrotemporal input patterns. These encoding neural populations would be subject to adaptation, etc, and could allow us to make predictions about the perception of novel stimuli.

1 underlying: 1.20 2.22, 1.20 2.22 5000

Our theoretical solution can be modified to account for experimental data in which the initial percept duration distribution is different from the steady state. However, there are circumstances in which inertia is fairly trivial, such as when buildup resets after a switch in attention (?). Stationary distributions might be appropriate for such circumstances. We can take a more abstract view and consider a steady-state buildup function constructed from averaging over timepoints aligned by switches into the grouped percept.

3.2 SWITCH-TRIGGERED AVERAGING PRODUCES STEADY STATE BUILDUP FUNCTION

To address the issue of inertia and the longer mean duration of the initial percept than subsequent grouped percepts on a trial, we considered a new method for constructing the buildup function: switch-triggered averaging. This method allows us to produce a

0.9 0.8

LSE: 1.36 1.95, 1.36 1.95 4000

3000

2000

true: 1.20, 2.22

fit: 1.36, 1.95

0.7 1000

0.6 0.5 0.4 0.3 0.2 0.1

00 5 10 15 20 25 duration(s)

5000 4000 3000 2000 1000

true: 1.20, 2.22

fit: 1.36, 1.95

00 5 10 15 00 5 10 15 20 25

Figure 6: The same as Figure ??, except that the Monte Carlo simulations have been constrained so that the gamma densities for each of the two states are matched. In addition, the least squares fit is constrained to find only two parameters. The recovered parameters for the duration distributions almost always match those used to generate the Monte Carlo simulated buildup function.

duration(s)

buildup function from a single long trial. Ignoring the first and second percept duration, we constructed buildup functions by estimating the probability over time for the split percept based on an event-triggered average aligned to each switch into the grouped percept. This method produces a buildup function at steady state, the probability of perceiving the split organization not just from the beginning of the trial but rather from the beginning of any grouped percept duration over the course of a long presentation.

3.3 DO PERCEPT-TO-PERCEPT CORRELATIONS MATTER FOR DESCRIBING BUILDUP IN AN ALTERNATING RENEWAL PROCESS?

The previous competition model simulations provide a test-bed for our novel statistical model. When given dominance durations are not statistically independent, and there is history dependence between successive

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perceptual epochs, can the alternating renewal process theory still relate buildup to the underlying duration distributions? We measured the correlations in the data produced for noise-driven and adaptation-driven alternations. As previously described, the adaptation-driven perceptual timecourses showed moderate correlations between the durations of successive percepts. Our statistical model, however, ignores this history dependence entirely, treating percept durations as independent random variables described by stationary density functions.

Importantly, even for adaptation-driven switching, for which history dependence between successive durations is prominent, the buildup function predicted by the alternating renewal process without any consideration for correlations matched the competition model simulated buildup function. We believe that the failure of history dependence to affect the dynamics of buildup results from the loss of information about individual trial timecourses caused by averaging. While the correlations between perceptual epochs may be important for understanding the timecourse of a particular trial, the correlations between specific perceptual epochs are ”washed out” by taking the average of many timecourses. Therefore, the dynamics of buildup can be described sufficiently well by the underlying distributions of dominance durations.

Previous accounts of buildup (??) have pointed to the accumulation of adaptation as a critical feature for the switch from a grouped to a split percept. Indeed, multi-second habituation in the auditory periphery (?) can predict the behavioral buildup of streaming. It may therefore be surprising that the alternating renewal process neglects to account for adaptation entirely. We believe that previous observations and our own can be reconciled, and may even be complementary. Previous work provides some account for how the switch from the grouped to the split percept might be accomplished as a result of an accumulation of adaptation. While we do not explicity invoke an accumulative process to describe this switch, the use of a gamma density to describe the distribution of dominance durations for the grouped percept implies history dependence. This is because the hazard function for a gamma density, in contrast to an exponential distribution, is dependent on time elapsed, evolving from 0 at time zero to a steady state value. The time dependence of the probability of switching out of the grouped percept may therefore be complementary with previous descriptions. What is still missing from these theories, however, is the mechanism by which the perceptual state switches back and undergoes alternations; how do we account for switches out of the split percept, and the distribution of split durations?

The renewal process model is agnostic to the specific mechanism by which states are found and alternations occur; we have used existing competition models for the sake of illustration and as a computational test-bed. Similar competition-like processes have been used to explain the alternations observed in ambiguous motion and stream segregation (? experiments). The specific mechanisms for switching in human stream segregation may be more complex; (?) find that feedforward and feedback processes in a thalamocortical loop might be differentially engaged for switches into and out of the perceptual organization that is strongest.

single switch case is just a distribution of integrated percept/first switch times and a distribution of split percept durations that is centered at a mean far exceeding the time course of data collection. sequence detector our statistical model is good

3.4 SUMMARY

We have presented a statistical model that explains how the distribution of state durations for an alternating renewal process at steady state describes the system’s behavior before reaching steady state.

3.5 DATA SHARING

Frontiers supports the policy of data sharing, and authors are advised to make freely available any materials and information described in their article, and any data relevant to the article (while not compromising confidentiality in the context of human-subject research) that may be reasonably requested

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by others for the purpose of academic and non-commercial research. In regards to deposition of data and data sharing through databases, Frontiers urges authors to comply with the current best practices within their discipline.

4 MATERIAL & METHODS

4.1 DERIVATION FOR ALTERNATING RENEWAL PROCESS

In the alternating renewal process there are 2 random variables, S(t) and Z(t). S(t) is the random elapsed time since last switching into the current state, evaluated at time t. Z(t) is a dichotomous random variable, where Z(t) in 0, 1 codes for the percept, grouped or split. For sake of convenience, we introduce 2 probability density -mass functions: f0(s, t)ds = P r{S(t)} in (s, s + dt) and Z(t) = 0, and f1(s, t)ds = P r(S(t) in (s, s + ds) and Z(t) = 1. The coupled pair of partial differential equations describing the evolution over time of these 2 probability density-mass functions is as follows:

∂f0(s,t) = − ∂ (1∗f0(s,t))−hT0(s)f0(s,t) (1) ∂t ∂s

∂f1(s,t) = − ∂ (1∗f1(s,t))−hT1(s)f1(s,t) (2) ∂t ∂s

where s is the elapsed time since entering state i and hTi (s) is the hazard function, the probability per unit time of exiting the current state at time t. hTi (s) = fTi , the ratio of the density function of durations

FˆTi Ti for state i and its complementary cumulative distribution function.

The value of S(t) is reset to 0 whenever an alternation between states occurs. The initial flux of probability (a source) at s = 0 is determined by the probability of switches out of the previous state, leading to the following boundary conditions:

f0(0, t) =

f1(0, t) =

∫∞ 0

∫∞ 0

hT1 (s)f1(s, t)ds (3) hT0 (s)f0(s, t)ds (4)

The probability that Z(t) = 1,p1(t), is the marginal probability mass function evaluated at z = 1. It is obtained by integrating f1(s,t) over all s; and similarly for Z(t) = 0. We used the following initial conditions, correponding to Pr{Z(t = 0) = 1} = 1, and Pr{Z(t = 0) = 0} = 0. In order to allow for a different distribution for the initial percept, we first solve for the case of beginning in state 1:

f1(s,t = 0) = δ(s);f0(s,t = 0) = 0 (5) For sake of simplifying notation, we define p1|1(t) = Pr{Z(t) = 1|Z(t = 0) = 1}. Using these

conditions yields the following solution:

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p1|1(t) = FT1(t) + 2π where f(x) is the Fourier transform of f(x)

ˆ

∫∞[] ̃ 1 1 ˆ ̃ ˆ ˆ iω iωt

−∞

dwiω FT1(ω)fT0(ω)fT1(ω) ˆ ˆ e (6) 1 − fT0 (ω)fT1 (ω)

To find the probability that Z(t) = 1, given Z(t = 0) = 0, p1|0, we time shift the above expression by the durations of the initial state, T0, whose density function is pT0 (t) [changed f to p]. This amounts to a convolution of the density for initial durations (and first switch times) with the previous solution:

[again changed f to p]

the same density function as all other T0, we find:

p1|0(t) = Thus the solution can be given in the Fourier domain as:

∫t 0

fT0 (s)f1([t − s]|z(0) = 1)ds (7) pˆ (ω)=fˆ0(ω)pˆ (ω) (8)

1|0 T0 1|1 Using the simplifying assumption that fT0 (t) = fT0 (t), that is, that the initial percept duration is from

( fˆ2 (ω)fˆ (ω) ) ˆ ˆ ̃ ˆ T0 T1

f1|0(ω) = FT1 (ω) fT0 + 1 − fˆ (ω)fˆ (ω) (9) T0 T1

ˆ At ω = 0, f (ω) = μ(T1)/(μ(T1) + mu(T0). From here, the expression in the time domain is obtained by

taking the inverse Fourier transform and finding the integral from 0 to t.

COMPETITION MODEL SIMULATIONS

Competition model simulations followed the procedures reported previously in ? for population firing rate model with spike frequency adaptation. Specifically,

= −u1 +f(−βu2 −γa1 +I1 +n1) =−a1+u1

 u ̇ 1  τ a a ̇ 1

 n ̇1

√

=−n1+σ2η(t) τn τn

= −u2 +f(−βu1 −γa2 +I2 +n2) =−a2+u2

 u ̇ 2  τ a a ̇ 2

 n ̇2

√

=−n2+σ2η(t) τn τn

The variable u1 corresponding to the short-time averaged firing rate of the population representing the “grouped” perceptual state, and u2 the firing rate of the population representing the “split” perceptual state. The variables a1 and a2 represent the spike-frequency adaptation. Parameter γ controls the strength of the adaptation, and β controls the strength of suppression from the competing population. I1 and I2 are the external inputs driving the two populations, and n1 and n2 are independent Ornstein-Uhlenback noise generators with mean zero and variance σ, and a timescale of τn. The input-output function used was a sigmoid, with f(x) = 1/(1 + exp((xθ)/k)).

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The simulation was carried out in nondimensionalized time, with the convention that one unit of time

corresponds to 10 msec. Time constants given in simulation time units were τa = 200, τn = 10. The

following parameter values are used: k = 0.1, θ = 0, β = 1. The values of the external inputs to the

populations I and I , the adaptation gain γ and the noise strength σ were varied as specified in the

√

figures, with the value of σ scaled in relation to the integration time step by 1/ dt to keep specified variance per unit time. Simulations were implemented in MATLAB using forward Euler integration with a time step of 0.1 (1 msec real time).

For each combination of parameter values, we simulated 500 trials of length 10 s with initial conditions u2(0), a1(0), a2(0), n1(0), n2(0) = 0 and u1(0) = 0.5; thus, at the beginning of each simulated trial, the first population to become dominant was always that corresonding to the first percept. With the resulting population firing rate timecourses, we obtained dominance durations by finding time points of the zero crossings of the differences of the firing rates. Using the samples of dominance durations obtained for each population (over 1000 durations for each population with each parameter set), we fitted gamma densities using maximum likelihood estimation. Simulated experimental buildup curves were constructed by computing the average for each timepoint across trials of the binary timecourse u2 > u1.

4.2 MONTE CARLO SIMULATIONS

We used Monte Carlo simulations to test the analytical solution above, and to evaluate its performance in relating the buildup function to the underlying distributions of dominance durations.

To generate a wide range of plausible buildup functions, we simply specify two distributions with random parameters within the bounds [1,5]. These were decided upon arbitrarily after visual inspection of many Monte Carlo simulated buildup functions. One of these distributions is labelled as the grouped percept durations, and the other as the split percept durations. For a given simulated trial timecourse, we draw alternating random samples from each of these two distributions. We begin with the distribution corresponding to the grouped state, and continue drawing samples until the sum of all the durations exceeds the length of a trial. These trial durations are converted into discretized timecourses by assigning a value of 0 to time intervals during which the state corresponds to a grouped percept, and assigning a value of 1 to the time intervals for which the percept state was split. In Monte Carlo simulations, we produce 1000 such trial timecourses, then take the average at each time point.

To estimate the gamma parameters from these Monte Carlo generated buildup functions, we used the analytical solution (above) and searched for the 4 parameters that minimized the squared error between the analytical and the Monte Carlo generated buildup function.

All simulations were implemented in MATLAB. Maximum likelihood estimation was carried out using the gamfit function from the stats toolbox.

DISCLOSURE/CONFLICT-OF-INTEREST STATEMENT

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

ACKNOWLEDGEMENT

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Funding: Text Text Text Text Text Text Text Text.

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SUPPLEMENTAL DATA

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